The Age Of Globular Clusters In Light Of Hipparcos: Resolving the Age Problem?

Brian Chaboyer^{1,2}, P. Demarque³, Peter J. Kernan⁵ and Lawrence M. Krauss^{4,5,6}

ABSTRACT

We review five independent techniques which are used to set the distance scale to globular clusters, including subdwarf main sequence fitting utilizing the recent Hipparcos parallax catalogue. These data *together* all indicate that globular clusters are farther away than previously believed, implying a reduction in age estimates. We now adopt a best fit value $M_V(RR) = 0.39 \pm 0.08(stat)$ at $[Fe/H] = -1.9$ with an additional uniform systematic uncertainty of $^{+0.13}_{-0.18}$. This new distance scale estimate is combined with a detailed numerical Monte Carlo study (previously reported by Chaboyer *et al.* 1996a) designed to assess the uncertainty associated with the theoretical age-turnoff luminosity relationship in order to estimate both the absolute age and uncertainty in age of the oldest globular clusters.

Our best estimate for the mean age of the oldest globular clusters is now 11.5 ± 1.3 Gyr, with a one-sided, 95% confidence level lower limit of 9.5 Gyr. This represents a systematic shift of over 2σ compared to our earlier estimate, due completely to the new distance scale—which we emphasize is not just due to the Hipparcos data. This now provides a lower limit on the age of the universe which is consistent with either an open universe, or a flat, matter dominated universe (the latter requiring $H_0 \leq 67 \text{ km s}^{-1} \text{ Mpc}^{-1}$). Our new study also explicitly quantifies how remaining uncertainties in the distance scale and stellar evolution models translate into uncertainties in the

¹Steward Observatory, University of Arizona, Tucson, AZ 85721 E-Mail: chaboyer@as.arizona.edu

²Hubble Fellow

³Department of Astronomy, and Center for Solar and Space Research, Yale University, Box 208101, New Haven, CT 06520-8101 E-Mail: demarque@astro.yale.edu

⁴Theory Group, CERN, CH-1211 Geneva, Switzerland

⁵Department of Physics, Case Western Reserve University, 10900 Euclid Ave., Cleveland, OH 44106-7079 E-Mail: pete@theory2.phys.cwru.edu; krauss@theory1.phys.cwru.edu

⁶Also Department of Astronomy

derived globular cluster ages. Simple formulae are provided which can be used to update our age estimate as improved determinations for various quantities become available. Formulae are also provided which can be used to derive the age and its uncertainty for a globular cluster, given the absolute magnitude of the turn-off, or the point on the subgiant branch 0.05 mag redder than the turn-off.

Subject headings: stars: interiors – stars: evolution – stars: Population II – globular clusters: general – cosmology: theory– distance scale

1. Introduction

The absolute age of the oldest Galactic globular clusters (GCs) currently provides the most stringent lower limit to the age of the universe, and as such, provides a fundamental constraint on cosmological models. In particular, for some time the best GC age estimates have been in direct contradiction with the maximum Hubble age for the preferred cosmological model, a flat matter dominated universe. The most recent comprehensive analyses suggested a lower limit of approximately 12 Gyr for the oldest GC's in our galaxy (e.g. Chaboyer, Demarque, Kernan & Krauss 1996, hereafter Paper I), which, for a flat matter dominated model, implies $H_0 \leq 53 \text{ km s}^{-1} \text{ Mpc}^{-1}$, a value which is low compared to almost all observational estimates.

Because of this apparent discrepancy, it remains critically important to continue to re-evaluate the errors associated with the GC age determination process itself. GC age estimates are obtained by comparing the results of theoretical stellar evolution calculations to observed color magnitude diagrams. The absolute magnitude of the main-sequence turn-off ($M_v(\text{TO})$) has small theoretical errors, and is the preferred method for obtaining the absolute ages of GCs (e.g. Renzini 1991). Age determination methods which utilize the color of the models, or post main-sequence evolutionary models are subject to much larger theoretical uncertainties, and do not lead to stringent age limits.

In recent years, a number of authors have examined the question of the absolute age GCs (e.g. Chaboyer & Kim 1995, Mazzitelli *et al.* 1995, Salaris *et al.* 1997) using different assumptions for the best available input physics. Chaboyer 1995 presented a table of absolute GC ages based on a variety of assumptions for the input physics needed to construct the theoretical age- $M_v(\text{TO})$ relationship. Vandenberg *et al.* 1996 have presented a review of the absolute ages of the GCs, and by comparing results from different authors, include an discussion on how various uncertainties in the age dating process effect the final age estimate.

In order to obtain both a best estimate, and a well-defined lower limit to the absolute age of the oldest GCs, we earlier adopted a direct approach of running a Monte Carlo simulation. In our Monte Carlo, the various inputs into the stellar evolution codes were varied within their inferred uncertainties, utilizing 1000 sets of isochrones, and the construction of over 4 million stellar models. From these theoretical isochrones, the age- $M_v(\text{TO})$ relationship was determined, and combined

with an empirical calibration of $M_V(\text{RR})$ ⁷ in order to calibrate age as a function of the difference in magnitude between the main sequence turn-off and horizontal branch ($\Delta V_{\text{HB}}^{\text{TO}}$). This calibration was used to derive the mean age of 17 old, metal-poor GCs using $\Delta V_{\text{HB}}^{\text{TO}}$. The principal result of this work was an estimate for the age of the oldest GCs of 14.6 ± 1.7 Gyr, with the one-sided 95% C.L. lower bound of 12.1 Gyr (Paper I) mentioned above. Another important result was an explicit demonstration that the uncertainty in $M_V(\text{RR})$ overwhelmingly dominated the uncertainty in the GC age determination. We chose a Gaussian distribution for the uncertainty in $M_V(\text{RR})$ because the data, while scattered, appeared to be appropriately distributed about the mean value, which we then determined to be $M_V(\text{RR}) = 0.60$ at $[\text{Fe}/\text{H}] = -1.9$, with an uncertainty of approximately 0.16 at the 95% confidence level.

Since this work was completed, the Hipparcos satellite has provided improved parallaxes for a number of nearby subdwarfs (metal-poor stars) (Perryman *et al.* 1997), the distance to a GC has been estimated using white dwarf sequence fitting (Renzini *et al.* 1996), a number of new astrometric distances to GCs have been published (Rees 1996), and improved theoretical horizontal branch models have become available (Demarque *et al.* 1997). This has lead us to critically re-evaluate the globular cluster distance scale (and hence, the $M_V(\text{RR})$ calibration), and update our estimate for the absolute age of the oldest GCs. We find, using the full Hipparcos catalogue along with the other independent distance estimators that all the data suggests that this distance scale, and hence the GC age estimate have shifted by a significant amount, suggesting that the dominant uncertainty in $M_V(\text{RR})$ was, and still is, not statistical but rather systematic in character.

A detailed discussion of the globular cluster distance scale is presented in §2. The input parameters and distributions in the Monte Carlo are presented in §3. The principal results of this paper are presented in §4, which includes simple formulae which can be used to update the absolute age of the oldest GCs when improved estimates for the various input parameters become available. Finally, §5 contains a brief summary of our results, and a brief discussion of their cosmological implications.

2. The Globular Cluster Distance Scale

It is currently impossible to directly determine distances to GCs using trigometric parallaxes. While such distance estimates may be available in the future from microarcsecond space astrometry missions (Lindegren & Perryman 1996, Unwin, Boden & Shao 1996), at present a variety of secondary distance estimates are all that is available for GC's. The different techniques rely upon different data and assumptions. As such, we have elected to review a number of these techniques, and present a GC distance scale which is based on combining 5 independent estimates. To facilitate

⁷Throughout this paper, the terms $M_V(\text{RR})$ ($V(\text{RR})$) and $M_V(\text{HB})$ ($V(\text{HB})$) will be used interchangeably, with the understanding that $M_V(\text{HB})$ ($V(\text{HB})$) refers to the mean magnitude of the HB in the instability strip.

this, we have reduced the various distance estimates to a calibration of $M_V(\text{RR})$. This allows us to derive GC ages via the $\Delta V_{\text{HB}}^{\text{TO}}$ method (see §4). As we are interested in absolute ages, we have focused our attention on those techniques which rely upon the minimum number of assumptions and thus hopefully should provide *a priori* the most reliable absolute distances.

2.1. Astrometric Distances

A comparison of the proper motion and radial velocity dispersions within a cluster allows for a direct determination of GC distances, independent of reddening (Cudworth 1979). Although this method requires that a dynamical model of a cluster be constructed, it is the only method considered here which directly measures the distance to a GC without the use of a ‘standard’ candle. The chief disadvantage of this technique is its relatively low precision. This problem is avoided by averaging together the astrometric distances to a number of different GCs. Rees (1996) presents new astrometric distances to eight GCs, along with two previous determinations. As pointed out by Rees, there are possibly large systematic errors in the dynamical modeling of M15, NGC 6397 and 47 Tuc. As such, these clusters will be excluded in our analysis. In addition Rees (private communication) cautions that the distance to M2 will be revised soon to due to a new reduction of the M2 proper motions. Excluding this cluster from the analysis results in six clusters whose distances have been estimated astrometrically. Table 1 tabulates the astrometric distances from Rees (1996). Unless otherwise noted, the numbers are those given by Rees (1996).. For the $[\text{Fe}/\text{H}]$ values, we have given preference to the high dispersion results of Kraft, Sneden and collaborators. Taking the weighted average of the $M_V(\text{RR})$ values listed in Table 1 results in $M_V(\text{RR}) = 0.59 \pm 0.11$ at $< [\text{Fe}/\text{H}] > = -1.59$, where the average $[\text{Fe}/\text{H}]$ value has been calculated using the same weights as in the $M_V(\text{RR})$ average.

2.2. White Dwarf Sequence Fitting

Renzini *et al.* (1996) have utilized deep HST WFPC2 observations of NGC 6752 to obtain accurate photometry of the cluster white dwarfs. They have combined this with similar observations of local white dwarfs with known parallaxes and masses (close to those in the cluster) to derive the distance to NGC 6752 using a procedure similar to main sequence fitting. The derived distance modulus is $(m - M)_O = 13.05 \pm 0.10$ assuming $E(B - V) = 0.04$. This reddening estimate is from Zinn 1985, and is identical to those found by Burnstein & Heiles 1982 and Carney 1979. NGC 6752 is a moderately metal-poor ($[\text{Fe}/\text{H}] = -1.54$ from Zinn & West 1984) cluster with $V(\text{HB}) = 13.63$ (tabulated by Chaboyer *et al.* 1996c). NGC 6752 has an extremely blue HB, thus, an estimate of $V(\text{HB})$ relies upon an extrapolation of the observed photometry. As such the determination of $V(\text{HB})$ in NGC 6752 is rather uncertain, and so have elected to take a rather generous error bar in the determination of $V(\text{HB})$ of ± 0.1 mag. Combining the above quantities yields $M_V(\text{RR}) = 0.45 \pm 0.14$.

2.3. Subdwarf Main Sequence Fitting

Using parallaxes of nearby field stars, it is possible to define the position of the ZAMS, and via a comparison to deep GC color magnitude diagrams obtain a rather direct estimate of the distance to a cluster. Unfortunately, the position of the ZAMS is a rather sensitive function of metallicity, and there are few nearby subdwarfs. Hence, there are few metal-poor stars with well determined absolute magnitudes.

The release of the Hipparcos data (Perryman *et al.* 1997) has improved this situation somewhat, providing a large database of high quality parallax measurements. The Hipparcos catalogue contains over 100,000 stars, of which nearly 21,000 stars have parallax errors less than 10%. The Hipparcos catalogue has been searched for stars which are suitable for GC main sequence fitting. When selecting stars for use in main sequence fitting, it is important to avoid potential biases due to unresolved binaries and stars which are evolved off the ZAMS. Known or suspected binaries which are not resolved photometrically should be avoided as both magnitudes and colors may be significantly altered by the presence of a companion. The use of stars which have evolved off the ZAMS may lead to systematic errors in the derived distance moduli, as it is not clear if GCs and metal-poor field stars are exactly of the same age. For example, a 2 Gyr age difference between a calibrating subdwarf at $M_V = 5$ and a GC would lead to a systematic error of 0.14 mag in the distance modulus (based on our standard isochrones). To be safe, we will only consider stars with $M_V \gtrsim 5.5$. Fainter than this, the stars are essentially unevolved.

Many of the stars in the Hipparcos catalogue have large relative parallax errors and are not useful for main sequence fitting. For these reason, we have elected to only consider stars with $\sigma_\pi/\pi < 0.10$. This stringent selection criterion was selected to minimize potential Lutz-Kelker type biases (Lutz & Kelker 1973, Brown *et al.* 1997). The Hipparcos catalogue was searched for stars which (a) have $\sigma_\pi/\pi < 0.10$, (b) are fainter than $M_V \simeq 5.5$, and (c) are not known or suspected Hipparcos binaries or variables. This resulted in a list of 2618 stars, of which the great majority have near solar metallicity. As we are interested in the most metal-poor globular clusters, we require stars with $[\text{Fe}/\text{H}] \lesssim -1.0$. To identify the metal-poor stars in the Hipparcos sample, we have cross identified the above Hipparcos subsample with a variety of $[\text{Fe}/\text{H}]$ catalogues: the 1996 high resolution spectroscopic catalogue Cayrel de Strobel *et al.* 1997, the Carney *et al.* 1994 catalogue, and the measurements reported by Gratton *et al.* 1997 and Pont *et al.* 1997. Stars with $[\text{Fe}/\text{H}] \lesssim -1.0$ were selected for further study. Carney *et al.* 1994 and Pont *et al.* 1997 obtained numerous radial velocity measurements which could be used to identify potential binaries. Gratton *et al.* 1997 have search for binaries based on an excess infrared flux. Any stars which were known or suspected binaries were removed from our final list. In total only 10 stars in the Hipparcos catalogue pass our stringent selection criterion ($\sigma_\pi/\pi < 0.10$, $M_V \gtrsim 5.5$, $[\text{Fe}/\text{H}] \lesssim -1.0$, and not known or suspected binaries). Given the small parallax errors in our sample, Lutz & Kelker (1973) type corrections are expected to be small (Brown *et al.* 1997). This issue is explored via a detailed Monte Carlo analysis in Appendix A, where it is concluded that the sample is indeed free from systematic biases, and so no corrections have been applied to the absolute magnitudes derived from

the Hipparcos parallax observations. Another indication that possible Lutz-Kelker type corrections are small for our sample is that the maximum σ_π/π value is 0.08, well below our threshold of 0.10. This makes it extremely unlikely that stars whose true parallax are systematically smaller than the observed parallax are preferentially included in our sample (the source of Lutz-Kelker type biases).

Our subdwarf sample only has two stars in common with the sample of Pont *et al.* 1997. This is because the Pont *et al.* 1997 sample only includes five stars on the ZAMS ($M_V \gtrsim 5.5$). As discussed above, even as small as a 2 Gyr age difference between evolved subdwarfs and the GC will lead to systematic errors in the distance modulus of ~ 0.14 mag. Of the unevolved stars in the Pont *et al.* 1997 sample, three are known binaries, which we do not use. Pont *et al.* 1997 apply an average binary correction of +0.375 mag to the 6 binaries in their total sample. The Poisson (root N) noise in this correction is ± 0.15 mag. If one has a large sample of binaries ($\gtrsim 30$), then the approach taken by Pont *et al.* 1997 to include average binary corrections is sound. However, given that the small number statistics in the present sample results in a very large error in the binary correction, we believe it is best not to use the binaries.

Theoretical models predict that the location of the ZAMS is a sensitive function of metallicity. Even with the Hipparcos data, the current observations are not accurate (or numerous) enough to empirically derive the ZAMS location as a function of metallicity. There are only four subdwarfs whose absolute magnitudes are known to within ± 0.1 mag. Unfortunately, the colors predicted by the models are still rather uncertain, and so we do not have a reliable calibration of how the location of the ZAMS changes as a function of metallicity. Thus one should ensure that the mean metallicity of subdwarf sample used in the main sequence fitting should be as close as possible to the metallicity of the GC. This requires accurate metallicity determinations.

We have searched the literature for abundance analyses, based upon high dispersion, high signal to noise spectrum. King (1997) has performed a detailed abundance analysis of HD 134439 and HD 134440 (a common proper motion pair). Rather surprisingly, King finds that the abundances of the α -capture elements are consistently some ~ 0.3 dex below the vast majority of metal-poor field stars, and those observed in GC giant stars. Due to their relatively high abundance, theoretical models predict that the α -capture elements play an important role in determining the position of a star in the color magnitude diagram. Given the peculiar abundances in these two stars, we have elected not to use them in main sequence fitting. The calibrating subdwarf data for the remaining 8 stars is presented in Table 2. Most of the data in Table 2 has been taken from the Hipparcos catalogue. The reddening estimates are from Carney *et al.* 1994 when available, or Pont *et al.* 1997. The $[\text{Fe}/\text{H}]$ values are discussed in detail below.

The $[\text{Fe}/\text{H}]$ abundance of HD 193901 has been determined by a number of groups. Recent values of $[\text{Fe}/\text{H}]$ are -1.00 (Carretta & Gratton 1997), -0.98 (Aker *et al.* 1994) and -1.22 (Tomkin *et al.* 1992). The higher abundances derived by Aker *et al.* 1994 and Gratton *et al.* 1997 may be largely due to the different effective temperature scales adopted by these authors compared to

Tomkin *et al.* 1992. This is still a matter of active debate, so we have elected to simply average the above $[\text{Fe}/\text{H}]$ values. In our main sequence fitting analysis (§2.3.1 and §2.3.2), we explore the consequences of the various $[\text{Fe}/\text{H}]$ values. HD 145417 has not been extensively studied, and the only spectroscopic metallicity determination is $[\text{Fe}/\text{H}] = -1.15 \pm 0.13$ (Gratton *et al.* 1997).

Balachandran & Carney (1996) have presented a detailed abundance analysis of HD 103095 (Groombridge 1830), the subdwarf with the best determined absolute magnitude. They found $[\text{Fe}/\text{H}] = -1.22 \pm 0.04$. Other abundances appeared to be typical of metal-poor stars. This is very similar to the value obtained by Gratton *et al.* 1997 ($[\text{Fe}/\text{H}] = -1.24 \pm 0.07$). The Balachandran & Carney (1996) value is adopted in this work. The abundance of HD 120559 has been determined to be $[\text{Fe}/\text{H}] = -1.23 \pm 0.07$ (Axer *et al.* 1994). Tomkin *et al.* 1992 have found $[\text{Fe}/\text{H}] = -1.45$ for HD 126681.

The metallicity of BD+59 2407 is $[\text{Fe}/\text{H}] = -1.60 \pm 0.16$ (Gratton *et al.* 1997) based on high signal to noise data. Carney *et al.* 1994 found $[\text{Fe}/\text{H}] = -1.91$ in their low signal to noise data. An examination of the Gratton *et al.* 1997 and Carney *et al.* 1994 abundances indicates that the former are systematically more metal-rich than the later. Gratton *et al.* (1997) found that on average, their $[\text{Fe}/\text{H}]$ values were +0.34 dex more metal rich than Tomkin *et al.* 1992. Once again, part of this difference is attributable to the different effective temperature scales. We have again elected to adopt the average value $[\text{Fe}/\text{H}] = -1.75$, and will discuss the different $[\text{Fe}/\text{H}]$ scales in §2.3.1 and §2.3.2.

A detailed abundance analysis of HD 25329 has been presented by Beveridge & Sneden (1994). They found $[\text{Fe}/\text{H}] = -1.84 \pm 0.05$, while Gratton *et al.* 1997 report $[\text{Fe}/\text{H}] = -1.69 \pm 0.07$. Beveridge & Sneden (1994) note that HD 25329 is a N-enhanced star; only $\sim 3\%$ of observed halo dwarfs are N-enhanced. However, they find that the relative abundances of α -capture and iron peak elements are normal for metal-poor stars. Nitrogen comprises only $\sim 3\%$ of the mass fraction of the heavy elements in a star. Thus, the fact that HD 25329 is N-enhanced is unlikely to affect its position on the color magnitude diagram, and so it will be used in the main sequence fitting. Taking a simple average of the above two $[\text{Fe}/\text{H}]$ determinations results in $[\text{Fe}/\text{H}] = -1.76$.

Finally, CPD-80 349 has $[\text{Fe}/\text{H}] = -2.26 \pm 0.2$ (Pont *et al.* 1997). This is based on low signal to noise spectra, and is on the same system as Carney *et al.* 1994. This is apparently the most metal-poor star in the Hipparcos catalogue (see Figure 1), and an improved abundance determination would be of great benefit. This star has $E(B - V) = 0.02$ (Pont *et al.* 1997). Given the poor quality of the abundance determination, and the fact that it is difficult to determine the reddening for a single star, we have elected not to use CPD-80 349 in our main sequence fits.

Figure 1 presents the HR diagram for the calibrating subdwarf data, along with a comparison with our standard isochrones. Most of the stars have metallicities in the range -1.1 to -1.5 and provide a nice calibration of the ZAMS in this $[\text{Fe}/\text{H}]$ range. The position of HD 25329 ($[\text{Fe}/\text{H}] = -1.76$) is somewhat surprising, as it lies along the same isochrone as HD 103095 ($[\text{Fe}/\text{H}] = -1.22$). Both of these stars have very well determined metallicities and parallaxes. Note that BD+59

2407 ($[\text{Fe}/\text{H}] = -1.75$) does not lie along the same isochrone as HD 25329. This could be due to an error in the reddening or parallax of BD+59 2407. Alternatively, it suggests that HD 25329 is anomalously bright for its metallicity and color. Clearly more data is needed to differentiate between these hypothesis. Unfortunately, an inspection of the Hipparcos catalogue reveals there are no candidate metal-poor, unevolved single stars with $\sigma_\pi/\pi < 0.10$ which are likely to have $[\text{Fe}/\text{H}] < -1.5$.

Given the $[\text{Fe}/\text{H}]$ values of the calibrating subdwarfs, accurate GC distances using main sequence fitting can be obtained for GC with $-1.8 \lesssim [\text{Fe}/\text{H}] \lesssim -1.1$. To perform main sequence fitting, accurate photometry well below the main sequence turn-off is required. In addition, the cluster $[\text{Fe}/\text{H}]$ value must be well determined. Given these restrictions, accurate main sequence fitting distances can only be determined to three globular clusters M5, M13 and NGC 6752. NGC 288 and 362 are not included, as moderate resolution spectra of a few cluster giants yield $[\text{Fe}/\text{H}]$ values which are considerably different (Carretta & Gratton 1997) from the Zinn & West (1984) values. Higher dispersion spectra of a number of stars in each cluster are required to determine accurate $[\text{Fe}/\text{H}]$ abundances for these two clusters. Our approach to subdwarf fitting differs significantly from that adopted by Reid (1997), Gratton *et al.* (1997) and Pont *et al.* (1997) in that (a) we do not use binary, evolved or chemically peculiar subdwarfs and (2) we do not make theoretical ‘color’ corrections to the subdwarf data to account for metallicity differences between the GC and subdwarfs, but ensure that the mean metallicity of the subdwarf sample is nearly identical to the GC. The subdwarf fitting results for each of the three clusters (NGC 6752, M5, M13) are discussed in turn.

2.3.1. NGC 6752

High resolution spectra of three giants yields $[\text{Fe}/\text{H}] = -1.58$ (Minniti *et al.* 1993), while the six giants studied by Norris & Da Costa 1995 yield $[\text{Fe}/\text{H}] = -1.52$. Carretta & Gratton 1997 obtained data for 4 other giants, and re-analyzed the above data to obtain $[\text{Fe}/\text{H}] = -1.42$. Averaging these three abundance determinations, we adopt $[\text{Fe}/\text{H}] = -1.51$. The reddening is $E(B - V) = 0.04$, as discussed in §2.2. Subdwarfs with $-1.23 \leq [\text{Fe}/\text{H}] \leq -1.76$ were used in the weighted, least squares fit. The mean abundance of these subdwarfs (using the same weighted as in the least squares fit to the NGC 6752 fiducial) is $[\text{Fe}/\text{H}] = -1.55$, very similar to our adopted $[\text{Fe}/\text{H}]$ abundance of NGC 6752.

The distance to NGC 6752 was determined using a weighted least squares fit to the deep photometry of this cluster as presented by Penny & Dickens 1986. The weights for the fit were the errors in the absolute magnitudes of the subdwarfs (σ_{M_V} , presented in Table 2). These absolute magnitude errors only include the parallax errors. To allow for errors in the photometry, an error of ± 0.02 mag was added in quadrature with the σ_{M_V} tabulated in Table 2 when performing the fit. The resultant distance modulus is $(m - M)_V = 13.33 \pm 0.04$ mag, where this error represents the error associated with the weighted least squares fit of the NGC 6752 fiducial to the subdwarf

data. To this error, one must add in errors associated with the reddening, and allow for possible metallicity errors. An uncertainty in the reddening of ± 0.01 translates into an error in the derived distance modulus of ± 0.05 .

Due to the possible systematic uncertainties in the metallicity abundances of the subdwarfs and NGC 6752 we have examined various possibilities, in order to determine how a possible mismatch between the metallicity of NGC 6752 and the calibrating subdwarfs might affect the distance modulus estimates.

1. Adopting the Carretta & Gratton 1997 abundance for NGC 6752 ($[\text{Fe}/\text{H}] = -1.42$), and using subdwarfs with $-1.22 \leq [\text{Fe}/\text{H}] \leq -1.76$. The weight of HD 103095 in the fit was decreased by increasing its σ_{M_V} error to ± 0.045 , ensuring that the mean weighted mean metallicity of the 5 calibrating subdwarfs was $[\text{Fe}/\text{H}] = -1.42$. The resultant distance modulus is $(m - M)_V = 13.31 \pm 0.03$ mag.
2. Adopting the Carretta & Gratton 1997 abundance for NGC 6752 along with the Gratton *et al.* 1997 and Axer *et al.* 1994 abundances for the subdwarfs. This was done as the Axer *et al.* 1994 and Gratton *et al.* 1997 subdwarf abundances are systematically more metal-rich than other determinations. HD 126681 does not have an abundance determinations by Gratton *et al.* 1997 or Axer *et al.* 1994 and was removed from the list. The 3 stars with $-1.24 \leq [\text{Fe}/\text{H}] \leq -1.69$ were used in the fit (HD 103095, 25329 and BD+59 2407). Once again, the weight of HD 103095 in the fit was decreased by increasing its σ_{M_V} error to ± 0.038 , ensuring that the mean weighted mean metallicity of the calibrating subdwarfs was $[\text{Fe}/\text{H}] = -1.42$. The resultant distance modulus is $(m - M)_V = 13.25 \pm 0.03$ mag.
3. Assuming that a systematic zero-point error (+0.20 dex) exists between the subdwarf $[\text{Fe}/\text{H}]$ determinations and NGC 6752 implying that NGC 6752 has a metallicity of $[\text{Fe}/\text{H}] = -1.71$ (in the subdwarf $[\text{Fe}/\text{H}]$ system). This results in the use of 3 stars with $-1.45 \leq [\text{Fe}/\text{H}] \leq -1.76$, and decreasing the weight of HD 25329 in the fit by increasing its σ_{M_V} error to ± 0.073 mag. The derived distance modulus is $(m - M)_V = 13.24 \pm 0.06$ mag.
4. Assuming that a systematic zero-point error (−0.20 dex) exists between the subdwarf $[\text{Fe}/\text{H}]$ determinations and NGC 6752 implying that NGC 6752 has a metallicity of $[\text{Fe}/\text{H}] = -1.31$ (in the subdwarf $[\text{Fe}/\text{H}]$ system). The three subdwarfs with $-1.22 \leq [\text{Fe}/\text{H}] \leq -1.45$ with equal weighting (implying a mean $[\text{Fe}/\text{H}] = -1.30$) were used in the fit. The resultant distance modulus is $(m - M)_V = 13.42 \pm 0.09$ mag.
5. Assuming the anomalous position of HD 25329 in the Figure 1 is due to an incorrect abundance determination, and so removing HD 25329 from the fit. The three subdwarfs with $-1.23 \leq [\text{Fe}/\text{H}] \leq -1.75$ were used in the fit, and the σ_{M_V} error in HD 120559 was increased to ± 0.22 so that the weighted mean subdwarf $[\text{Fe}/\text{H}]$ was -1.51 . The derived distance modulus is $(m - M)_V = 13.30 \pm 0.11$ mag.

6. Assuming that the $[\text{Fe}/\text{H}]$ determination of BD+59 2407 is in error, and removing it from the fit. This results in the use of the three subdwarfs with $-1.23 \leq [\text{Fe}/\text{H}] \leq -1.76$ and a distance modulus of $(m - M)_V = 13.34 \pm 0.04$ mag.

The maximum change in the derived distance modulus is ± 0.09 mag, which we take to be the $1\text{-}\sigma$ error in the distance modulus due to possible metallicity errors. Adding the metallicity, reddening and fitting errors together in quadrature yields a distance modulus of $(m - M)_V = 13.33 \pm 0.11$. As discussed in §2.2, $V(\text{HB}) = 13.63 \pm 0.1$, and so $M_V(\text{RR}) = 0.30 \pm 0.15$ from the subdwarf distance modulus. This subdwarf visual distance modulus corresponds to $(m - M)_O = 13.20 \pm 0.11$ (with $A_V = 3.2$), which is within $1\text{-}\sigma$ of the distance obtained from the white dwarfs $(m - M)_O = 13.05 \pm 0.10$ (Renzini *et al.* 1996).

2.3.2. M5

High dispersion spectroscopic analysis indicates that this cluster has $[\text{Fe}/\text{H}] = -1.17$ (Snedden *et al.* 1992). The reddening is $E(B - V) = 0.03$, as summarized by Reid 1997. A deep color magnitude diagram for this cluster has been presented by Sandquist *et al.* 1996. Subdwarfs with $-1.07 \leq [\text{Fe}/\text{H}] \leq -1.23$ were used in the weighted, least squares fit. The weighted mean abundance of these subdwarfs $[\text{Fe}/\text{H}] = -1.19$, very similar to our adopted $[\text{Fe}/\text{H}]$ abundance of NGC 6752. Subdwarf fitting yields a distance modulus of $(m - M)_V = 14.51 \pm 0.02$. To this error, one must add in errors associated with the reddening, and allow for possible metallicity errors. An uncertainty in the reddening of ± 0.01 translates into an error in the derived distance modulus of ± 0.05 .

Due to the possible systematic uncertainties in the metallicity abundances of the subdwarfs and M5, we have once again examined the effects that various scenarios for metallicity errors have on the derived distance modulus.

1. Adopting the Gratton *et al.* 1997 and Axer *et al.* 1994 abundances for the subdwarfs, and using the 4 subdwarfs with $-0.99 \leq [\text{Fe}/\text{H}] \leq -1.23$ (HD 103095, 120559, 145417 and 193901) resulting in a mean metallicity of the subdwarfs of $[\text{Fe}/\text{H}] = -1.20$ and $(m - M)_V = 14.51 \pm 0.03$.
2. Assuming that a systematic zero-point error ($+0.20$ dex) exists between the subdwarf $[\text{Fe}/\text{H}]$ determinations and M5 implying that M5 has a metallicity of $[\text{Fe}/\text{H}] = -1.37$ (in the subdwarf $[\text{Fe}/\text{H}]$ system). This results in the use of 5 stars with $-1.22 \leq [\text{Fe}/\text{H}] \leq -1.76$, and decreasing the weight of HD 103095 in the fit by increasing its σ_{M_V} error to ± 0.032 mag. The derived distance modulus is $(m - M)_V = 14.49 \pm 0.02$ mag.
3. Assuming that the $[\text{Fe}/\text{H}]$ value for HD 103095 is in error, and so removing it from the fit. This results in the use of three subdwarfs ($-1.07 \leq [\text{Fe}/\text{H}] \leq -1.23$ and $(m - M)_V = 14.58 \pm 0.03$ mag.

4. Removing HD 193901 from the fit (leaving HD 103095, 120559 and 145417 with a mean metallicity of $[\text{Fe}/\text{H}] = -1.20$). The resultant distance modulus is $(m - M)_V = 14.51 \pm 0.02$ mag.
5. Removing HD 145417 from the fit, and giving equal weight to the remaining three stars (HD 103095, 120559 and 193901) to ensure a mean $[\text{Fe}/\text{H}] = -1.17$. The derived distance modulus is $(m - M)_V = 14.58 \pm 0.07$ mag.
6. Only using HD 145417 ($[\text{Fe}/\text{H}] = -1.15$) in the fit, resulting in $(m - M)_V = 14.56 \pm 0.03$ mag.

The maximum change in the derived distance modulus is ± 0.07 mag, which we take to be the $1\text{-}\sigma$ error in the distance modulus due to possible metallicity errors. Adding the metallicity, reddening and fitting errors together in quadrature yields a distance modulus of $(m - M)_V = 14.51 \pm 0.09$. Utilizing $V(\text{RR}) = 15.05 \pm 0.02$ (Reid 1996), results in $M_V(\text{RR}) = 0.54 \pm 0.09$.

2.3.3. M13

High dispersion spectroscopic analysis indicates that this cluster has $[\text{Fe}/\text{H}] = -1.58$ (Kraft *et al.* 1997). The reddening is $E(B - V) = 0.02$ (Zinn & West 1984). A deep color magnitude diagram and fiducial has been obtained by Richer & Fahlman 1986. This is a somewhat difficult metallicity to deal with, as none of the calibrating subdwarfs has a metallicity near $[\text{Fe}/\text{H}] = -1.58$. We have explored a number of possible options for a subdwarf sample selection.

1. Utilizing the subdwarfs with $-1.23 \leq [\text{Fe}/\text{H}] \leq -1.76$, results in a weighted mean $[\text{Fe}/\text{H}] = -1.55$ for the subdwarfs and $(m - M)_V = 14.54 \pm 0.04$ is obtained from a weighted fit to the fiducial.
2. The subdwarf $[\text{Fe}/\text{H}]$ range is restricted to $-1.45 \leq [\text{Fe}/\text{H}] \leq -1.76$ and the stars are equally weighted (resulting in a mean $[\text{Fe}/\text{H}] = -1.65$) then the derived distance modulus is $(m - M)_V = 14.46 \pm 0.10$
3. A systematic offset error of -0.2 dex is assumed between the subdwarfs and M13, implying that M13 has $[\text{Fe}/\text{H}] = -1.78$ in the subdwarf system. In this case, only the two stars with $[\text{Fe}/\text{H}] = -1.75$ and -1.76 are used, resulting in $(m - M)_V = 14.39 \pm 0.05$
4. A systematic offset error of $+0.2$ dex is assumed between the subdwarfs and M13, implying that M13 has $[\text{Fe}/\text{H}] = -1.38$ in the subdwarf system. In this case subdwarfs with $-1.22 \leq [\text{Fe}/\text{H}] \leq -1.76$ are used, and the weight of HD 103095 in the fit is decreased by increasing its σ_{M_V} error to ± 0.034 (to ensure a subdwarf mean $[\text{Fe}/\text{H}] = -1.38$). The derived distance modulus is $(m - M)_V = 14.51 \pm 0.02$

The derived distance moduli vary from $14.39 - 14.54$. We have elected to adopt the mid-point as our best value, and utilize a generous $1\text{-}\sigma$ error of ± 0.09 , hence $(m - M)_V = 14.47 \pm 0.09$. Adding in quadrature the error due to reddening, a total distance modulus error of ± 0.10 is adopted. M13 has a very blue HB, so the determination of $V(\text{HB})$ is difficult. We adopt $V_{\text{HB}} = 14.83 \pm 0.10$ (Chaboyer *et al.* 1996c), implying $M_V(\text{RR}) = 0.36 \pm 0.14$.

2.4. Calibration of $M_V(\text{RR})$ via the LMC

Walker (1992) determined the mean magnitudes of a number of RR Lyr stars in several LMC clusters. Adopting a distance modulus to the LMC of $\mu_{\text{LMC}} = 18.50 \pm 0.10$, he found $M_V(\text{RR}) = 0.44 \pm 0.10$. This distance modulus was based upon the traditional calibration of Cepheids. Using Hipparcos based parallaxes, Feast & Catchpole (1997) derived $\mu_{\text{LMC}} = 18.70 \pm 0.10$. This distance relied upon a period-color relation, and parallaxes of rather low quality ($\sigma_\pi/\pi \gtrsim 0.3$). An analysis of the Hipparcos Cepheid data by Madore & Freedman 1997 yields $\mu_{\text{LMC}} = 18.57 \pm 0.11$ who noted that “other effects on the Cepheid PL relation (e.g. reddening, metallicity, statistical errors) are as significant as this reassessment of its zero point”. The distance to the LMC may be estimated independent of the Cepheid or RR Lyr⁸ from geometric considerations using the ‘light echo’ times to the ring around SN 1987A (Panagia *et al.* 1991, Gould 1995). Using the same data set, but independent analysis, the SN1987A ring distance has been re-calculated by a few groups. Sonneborn *et al.* 1997 found $\mu_{\text{LMC}} = 18.43 \pm 0.10$, while Gould & Uza (1997) determined $\mu_{\text{LMC}} < 18.44 \pm 0.05$. Recently, Lundquist & Sonneborn 1997 reported a lower limit of $\mu_{\text{LMC}} < 18.67 \pm 0.08$. In light of these contradictory results, we have elected to follow the conclusion of Madore & Freedman 1997 and adopt a distance modulus of 18.50 mag for the LMC, and assume an uncertainty ± 0.14 to fully encompass the range of recently published values. Adopting this distance modulus, along with the photometry of Walker (1992) yields $M_V(\text{RR}) = 0.44 \pm 0.14$ at $[\text{Fe}/\text{H}] = -1.9$.

2.5. Theoretical HB models

Continued advances in our understanding in the basic physics which governs stellar evolution have lead to ever more reliable theoretical HB models. Recently, Demarque *et al.* 1997 have constructed synthetic HB models for various clusters, based upon new evolutionary models for HB stars. Assuming a primordial helium abundance of 0.23, these models predict $M_V(\text{RR}) = 0.34$ for M92 ($[\text{Fe}/\text{H}] = -2.25$) and $M_V(\text{RR}) = 0.42$ for M15 ($[\text{Fe}/\text{H}] = -2.15$). These two clusters are among the 17 old clusters whose mean age is determined in §4. Taking a simple average of the above

⁸Reid (1997) and Gratton *et al.* 1997 have made an estimate of the LMC distance using his subdwarf fitting based on Hipparcos parallaxes, of $\mu_{\text{LMC}} \approx 18.65$ and $\mu_{\text{LMC}} \approx 18.60$. However, this is based in part on subdwarf fitting to M92 ($[\text{Fe}/\text{H}] = -2.2$). All of the $[\text{Fe}/\text{H}] < -2.0$ subdwarfs are either (a) evolved off the ZAMS, or (b) suspected/known binaries, or (c) have poor $[\text{Fe}/\text{H}]$ determinations.

two numbers yields $M_V(\text{RR}) = 0.38$ at $[\text{Fe}/\text{H}] = -2.20$. Also, since the primordial helium value utilized in this analysis was on the low side, we have adjusted this mean value to 0.36 to account for a mean primordial helium value of 0.235. We also adopt an error of 0.10 mag on $M_V(\text{RR})$, to allow for possible errors in the models and in the primordial helium abundance estimate.

2.6. Combining the Distance Estimates

The individual determinations of $M_V(\text{RR})$ at the various metallicities are summarized in Table 3. While it is not evident from this table, there is considerable evidence from other observations and theoretical modeling that $M_V(\text{RR})$ is a function of $[\text{Fe}/\text{H}]$:

$$M_V(\text{RR}) = \mu([\text{Fe}/\text{H}] + 1.9) + \gamma. \quad (1)$$

We have chosen distance calibrations which yield reliable absolute numbers with the minimum possible systematic uncertainties⁹ Hence, they are useful in deriving the value of γ . However, these $M_V(\text{RR})$ determinations do not provide reliable information on the $M_V(\text{RR})$ - $[\text{Fe}/\text{H}]$ slope μ . For this, one needs to utilize techniques which yield reliable relative $M_V(\text{RR})$ values as a function of $[\text{Fe}/\text{H}]$. Theoretical HB models, as well as Baade-Wesselink studies of field RR Lyr stars provide the best estimate of the $M_V(\text{RR})$ - $[\text{Fe}/\text{H}]$ relationship¹⁰ The semi-empirical Baade-Wesselink method has been applied by Jones *et al.* 1992 and Skillen *et al.* 1993. Reanalysis of these data suggest that $\mu = 0.22 \pm 0.05$ (Sarajedini *et al.* 1997). The latest theoretical models of blue HB clusters yield slopes of $\mu = 0.25 \pm 0.07$ (Demarque *et al.* 1997). A weighted mean value of $\mu = 0.23 \pm 0.04$ was adopted. The third column in Table 3 lists the various values of γ implied by the individual determinations of $M_V(\text{RR})$. There is a considerable spread in these values (0.21 to 0.52), reinforcing the notion that the dominant uncertainty remains systematic and that our previous procedure of assigning a Gaussian uncertainty to this quantity was ill-advised. We have thus now chosen to utilize a uniform top-hat uniform distribution which evenly weights all values in the range 0.21 to

⁹In contrast to Paper I, we have elected not to include the statistical parallax results of Layden *et al.* 1996. This technique yields a reliable estimate of γ for field RR Lyr stars. However, we now recognize that this number may be biased compared to globular cluster RR Lyr stars. Field RR Lyr stars are preferentially found in places where the evolutionary time-scales are long (ie: near the zero-age horizontal branch). No such selection effect exists for cluster RR Lyr stars, so one would suspect that the typical cluster RR Lyr stars would be more evolved, and hence, brighter than the typical field RR Lyr star. Indeed, the statistical parallax results lead to $\gamma = 0.62 \pm 0.12$, which is 0.2 mag fainter than what we determine for cluster RR Lyr stars.

¹⁰The synthetic HB models confirm the earlier result of Lee 1991 that $M_V(\text{RR})$ depends both on $[\text{Fe}/\text{H}]$ and HB morphology, a plausible explanation for the different slopes derived by different authors. For example, if one chooses the same group of globular clusters as Gratton *et al.* 1997, who derived a slope of 0.22 ± 0.09 by fitting Hipparcos subdwarfs to globular cluster main sequences, the theoretical models yield a slope near 0.25. The theoretical ZAMS slope is 0.20. Note that the models show that the evolutionary correction from the ZAMS used by Carney *et al.* 1992 should be used with caution, as it clearly does not apply to the more extreme HB morphologies observed in the oldest clusters.

0.52. However to give some emphasis to the mean value of the measured data, we have added to this distribution a Gaussian distribution centered on the weighted mean of $\gamma = 0.39$ with an uncertainty of 0.08, doubling the calculated error in the mean to account for the average deviation from the mean. Note that this new mean value is 0.21 mag (more than $2 - \sigma$) below the value adopted in Paper I. This will lead to a considerably downward revision in our GC age estimates, which we believe will also now have a distribution which is more appropriate to the systematic nature of the existing uncertainties. (We emphasize once again that while the Hipparcos data provided a motivation for re-examining this value, all of the other distance estimators we have examined appear now to be consistent, within the systematic uncertainties quoted, with a much lower value than we previously adopted.)

Spectroscopic studies of blue horizontal-branch (BHB) stars provide further support for the longer GC distance scale adopted here. From both the continuous spectrum and absorption line profiles, it is possible, with the help of model stellar atmospheres, to derive the effective temperature and surface gravities of these stars. This combined information yields the mass-to-light ratio M/L of the star, and if its distance is also known, its mass. This method has in the past yielded masses incompatible with the standard HB evolution theory (masses lower than evolutionary models) (de Boer *et al.* 1995; Moehler *et al.* 1995, 1997). A recent attempt to rederive the distances of some field BHB stars using Hipparcos parallaxes, could not be given much weight in view of the smallness of the parallaxes (de Boer *et al.* 1997). Using Reid’s (1997) reanalysis of the distances to some globular clusters based on larger Hipparcos subdwarf parallaxes, Heber *et al.* 1997 have reconsidered this problem, and concluded that the higher luminosities for BHB stars now yield masses in better agreement with the evolutionary masses. This important result provides independent support, based on physical modeling, for revising upward the distance scale to globular clusters, as suggested by several lines of reasoning, including the Hipparcos parallax data.

Finally, we should point out that the new distance scale yields a poor fit to calculated isochrones near the cluster turnoff. This suggests that the stellar model radii may need revision (better stellar atmospheres and convection modeling). Improvements in atmosphere models may lead to revisions of the T_{eff} to color transformations, particularly for the most metal poor stars. Since the ΔV_{HB}^{TO} method is little affected by surface effects, it further justifies our preference for this approach over the $\Delta(B - V)$ approach, and the approach of fitting to the shapes of theoretical isochrone turnoffs, both of which are sensitively affected by atmosphere and outer envelope physics.

3. The Monte Carlo Variables

In order to access the range of error associated with stellar evolution calculations and age determinations, the various inputs into the stellar evolution codes were varied within their uncertainties. In this Monte Carlo analysis, the input parameters were selected randomly from a given distribution. The distributions are based on a careful analysis of the recent literature, as summarized in Paper I. As ages will be derived using $M_V(TO)$ most attention was paid to parameters

which could effect the age- $M_V(\text{TO})$ relationship. Table 4 provides an outline of the various input parameters and their distribution. If the distribution is given as statistical (stat.), then the parameter in question was drawn from a Gaussian with the stated σ . If the error was determined to be a possible systematic (syst.) one, then the parameter was drawn from a top-hat (uniform) distribution. In total, 1000 independent sets of isochrones were calculated. Each set of isochrones consisted of three different metallicities ($[\text{Fe}/\text{H}] = -2.5, -2.0$ and -1.5) at 15 different ages (8 – 22 Gyr) (see Paper I for further details).

4. Results

4.1. The Technique

The absolute magnitude of the main sequence turn-off is the favored age determination technique when absolute stellar ages are of interest (see discussion in Paper I). Turn-off luminosity ages can be determined independent of reddening by using the difference in magnitude between the main sequence turn-off and the HB, $\Delta V_{\text{HB}}^{\text{TO}}$. Each set of Monte Carlo isochrones provides an independent calibration of $M_V(\text{TO})$ as a function of age. This was combined with the $M_V(\text{RR})$ calibration discussed in §2 to determine a grid of predicted $\Delta V_{\text{HB}}^{\text{TO}}$ values as a function of age and $[\text{Fe}/\text{H}]$ which is then fit to an equation of the form

$$t_9 = \beta_0 + \beta_1 \Delta V + \beta_2 \Delta V^2 + \beta_3 [\text{Fe}/\text{H}] + \beta_4 [\text{Fe}/\text{H}]^2 + \beta_5 \Delta V [\text{Fe}/\text{H}], \quad (2)$$

where t_9 is the age in Gyr. The observed values of $\Delta V_{\text{HB}}^{\text{TO}}$ and $[\text{Fe}/\text{H}]$, along with their corresponding errors, are input in (2) to determine the age and its error for each GC in our sample.

The age determination for any individual globular cluster has a large uncertainty, due to the large observational errors in $V(\text{TO})$. This error is minimized by determining the mean age of a number of globular clusters. However, there is a significant age range among the globular clusters (e.g. Sarajedini & Demarque 1990, VandenBerg *et al.* 1990, Buonanno *et al.* 1994b, Chaboyer *et al.* 1996c). This problem was avoided by selecting a sample of globular clusters which are well observed, metal-poor ($[\text{Fe}/\text{H}] \leq -1.6$), and which are not known to be young (based on HB morphology and/or the difference in color between the turn-off and giant branch). In the tabulation of Chaboyer *et al.* 1996c, 17 GCs satisfy the above criteria: NGC 1904, 2298, 5024, 5053, 5466, 5897, 6101, 6205, 6254, 6341, 6397, 6535, 6809, 7078, 7099, 7492, and Terzan 8. The observational data for each cluster was taken from Chaboyer *et al.* 1996c. The mean (and median) metallicity of this sample is $[\text{Fe}/\text{H}] = -1.9$.

4.2. A Likelihood Distribution for the Age of the Oldest Globular Clusters

To derive our best estimate for the age, and uncertainty in the age of the oldest GCs, a mean age and 1σ uncertainty in the mean was determined for each set of isochrones, and a given value of

$M_V(\text{RR})$. The value of $M_V(\text{RR})$ was taken to be a random variable, weighted as described earlier (a top hat distribution between 0.21 and 0.52 superimposed on a Gaussian distribution with mean and uncertainty $M_V(\text{RR}) = 0.39 \pm 0.08$), and the sets of isochrones were sampled with replacement 12,000 times. For each sample, we recorded a random age drawn from a Gaussian distribution with the mean age and variance for that isochrone set at that value of $M_V(\text{RR})$.

The age data were sorted and binned, to produce the histogram shown in Figure 2. The median and mean age is 11.5 Gyr, with a standard deviation of 1.3 Gyr. The 1-sided, 95% lower confidence limit is 9.5 Gyr, and is believed to represent a robust lower limit to the age of the GCs, and more properly takes into account the residual systematic uncertainties in $M_V(\text{RR})$, which largely determine the width of the derived age distribution. We are fully aware that due to our revision of the $M_V(\text{RR})$ zero-point, these ages are considerably reduced compared to the ages given in Paper I. Indeed, our new mean age is below our previous claimed 95% lower limit which was based on the assumption of Gaussian uncertainty in $M_V(\text{RR})$. In any case, our new results considerably alter the constraints one can derive on cosmological models (see §5).

Even though we have considered four independent distance determinations in addition to the Hipparcos parallaxes, our age estimate is in good agreement with two recent works which relied solely on Hipparcos parallaxes to determine the distances (and hence, ages) to a number of GCs (Reid 1997; Gratton *et al.* 1997). Pont *et al.* 1997 have determined an age of 14 Gyr for M92, which is in disagreement with our work. Pont *et al.* 1997 made a new fit of the CMD of M92 to theoretical isochrones, based on the Hipparcos subdwarf data. This paper represents a comprehensive analysis of the available data, and attempts the difficult task of correcting for selection effects which are more relevant in this case than the classical Lutz & Kelker 1973 corrections. However, our own analysis suggests that they have overestimated the corrections to the Hipparcos parallaxes due to biases (see Appendix). Their corrections due to the presence of binaries is very uncertain (± 0.15 mag); a fact which was not considered by Pont *et al.* 1997 in their analysis. A better procedure, which is not to include the suspected binaries in the fit, yields a larger distance modulus for M92. This approach, as pointed out by Pont *et al.* 1997 yields $(m - M)_V = 14.74 \pm 0.08$ mag. With this distance modulus, and the photometry of Stetson & Harris 1988 (the same photometry used by Pont *et al.* 1997), we calculate that the absolute magnitude of the point on the subgiant branch which is 0.05 mag redder than the turn-off is $V(\text{BTO}) = 3.39 \pm 0.08$ mag. This point is an excellent diagnostic of the absolute age of M92 (Chaboyer *et al.* 1996b), and using our isochrones (as outlined in §4.5) results in an age for M92 of 12.1 ± 1.3 Gyr. This is in good agreement with our estimate for the mean age of the oldest GCs (which includes M92) given above.

In the final analysis, the Pont *et al.* 1997 paper puts most of the weight of their fit on the agreement between the shapes of the theoretical isochrones and the data near the turnoff. However, this optimistic assessment of the models does not seem warranted in view of the well-known uncertainties associated with the treatment of convection, and the neglect of diffusion in the isochrones used (helioseismology has taught us that diffusion must be taken into account in the Sun (Basu *et al.* 1996, Guenther & Demarque 1997)). The need to apply an arbitrary color shift to the Vandenberg

et al. 1997 isochrones to reproduce the observed colors of M92, is another indication of the uncertainties involved, and lends further support to the choice of the $\Delta V_{\text{HB}}^{\text{TO}}$ method in dating globular clusters. We conclude that taking into account the differences in adopted distance moduli, and the neglect of diffusion by VandenBerg *et al.* 1997, our age estimate for M92, which is 11.5 ± 1.3 Gyr, is in good agreement with the Hipparcos data presented by Pont *et al.* 1997.

4.3. Effect of $M_V(\text{RR})$ on the age estimate

As was emphasized in Paper I (and by other authors), the principal uncertainty in absolute GC age determinations is the distance scale. With the $\Delta V_{\text{HB}}^{\text{TO}}$ age determination technique, this translates into the uncertainty in $M_V(\text{RR})$. We explicitly display this effect in Figure 3, where the GC ages are plotted as a function of $M_V(\text{RR})$. In order to quantify this uncertainty, median and $\pm 1\sigma$ points were determined as a function of $M_V(\text{RR})$. These were obtained by sorting the data based on $M_V(\text{RR})$, and then binning the ages as a function of $M_V(\text{RR})$. Sixty bins (corresponding to 200 ages per bin) were used, and in each bin the median age, and $\pm 1\sigma$ (68% range) ages were determined. An inspection of these points revealed that a simple linear relationship existed when one used the log of the age. A linear function of the form $\log(t_9) = a + b M_V(\text{RR})$ was fitted to this data, and the coefficients of this fit are given in the figure caption.

The median and $\pm 1\sigma$ fits are extremely useful summaries of our result. For example, at $M_V(\text{RR}) = 0.40$, the median fit yields 11.7 Gyr, identical to that given by the entire distribution (Fig. 2). The $\pm 1\sigma$ fits yield ages of 12.6 and 11.0 Gyr. Thus, if $M_V(\text{RR})$ was known to be exactly 0.40, then the error in the age of the oldest GCs would be ± 0.8 Gyr, due solely to the residual theoretical uncertainties in the stellar evolution calculations. The median and $\pm 1\sigma$ fits we present here may be used to update our age estimate as further data are obtained. For example, if $M_V(\text{RR}) = 0.50 \pm 0.05$ then from the fits, the median age would be 13.03 Gyr, with an error of ± 0.94 Gyr due to the theoretical uncertainties aside from those associated with $M_V(\text{RR})$. Next, from the median fit, the median age at $M_V(\text{RR}) = 0.45$ and 0.55 may be determined (corresponding to ± 0.05 mag) in order to estimate that the uncertainty in age associated with the $M_V(\text{RR})$ uncertainty is ± 0.68 Gyr. Combining these two error estimates in quadrature (± 0.94 and ± 0.68) would result in a best estimate of 13.1 ± 1.2 Gyr for $M_V(\text{RR}) = 0.50 \pm 0.05$. To verify this result, we have re-run the Monte Carlo analysis with the above choice of $M_V(\text{RR})$ and found identical results to those obtained from the $M_V(\text{RR})$ median and $\pm 1\sigma$ fits above.

4.4. Effect of the stellar evolution parameters on the age estimate

In order to examine how the individual stellar evolution parameters (given in Table 4) affect the estimated age, the mean age of the 17 GCs was determined for each of the 1000 Monte Carlo isochrones assuming fixed value $M_V(\text{RR}) = 0.40$. In a procedure analogous to that used for the

$M_V(\text{RR})$ fits, median and $\pm 1\sigma$ fits were determined for each of the 13 continuous variables listed in Table 4. As only 1000 points were available, only 20 bins were used. In addition, it was found that (due to the reduced age range), a linear fit provided as good a description as a log fit. Thus, the median and $\pm 1\sigma$ fits for each parameter x , were of the form $t_9 = a + bx$.

This procedure revealed that several of the input parameters had a negligible effect on the derived ages of the globular clusters. In order of importance, the following parameters were found to impact the GC age estimate: $\alpha/[\text{Fe}]$, mixing length, helium abundance, $^{14}\text{N} + \text{p} \rightarrow ^{15}\text{O} + \gamma$ reaction rate, helium diffusion coefficient, and the low temperature opacities. The plots of age as a function of these important parameters are shown in Figures 4 — 9. The figure captions give the coefficients of the median and $\pm 1\sigma$ fits for each of the variables. These fits can be used to update our best estimate for the age of the oldest globular clusters (in a manner analogous to that described for the $M_V(\text{RR})$ fits), as improved determinations of the above quantities become available.

In addition to the 13 continuous variables, we considered two binary variables (surface boundary condition, and color table, see Table 4). To examine the effect these parameters have on the derived ages, the ages were divided into 2 groups depending on which surface boundary condition (color table) was used in the stellar evolution codes. Histograms were constructed for each group, and compared. Not surprisingly, we found that the choice of the surface boundary condition had a negligible impact on the derived ages. However, the choice of the color table was important, and the two histograms are plotted in Figure 10. The choice of the color table changes the median age by 0.7 Gyr.

4.5. Calibration of the $M_V(\text{TO})$ and $M_V(\text{BTO})$ age relations

If the distance modulus to some cluster is known, then an accurate absolute age may be determined using $M_V(\text{TO})$ or alternatively using $M_V(\text{BTO})$ (Chaboyer *et al.* 1996b). This later point is defined to be the point on the subgiant branch which is 0.05 mag redder (in B—V) than the turn-off. As we have discussed, this point is easy to measure on an observed color magnitude diagram, yet has similar theoretical uncertainties to $M_V(\text{TO})$ (Chaboyer *et al.* 1996b). As a result, the precision in age estimation for individual clusters is better using $M_V(\text{BTO})$. The Monte Carlo isochrones may be used to quantify the error associated with an age determined via either method. To facilitate such error estimates, we have calculated the median and $\pm 1\sigma$ $M_V(\text{BTO})$ ($M_V(\text{TO})$) points as a function of age (in a manner similar to that described in the previous subsection) for four values of $[\text{Fe}/\text{H}]$: -2.5 -2.0 -1.5 and -1.0 . For ages between 8 and 17 Gyr, these points were then fit to a function of the form

$$\log(t_9) = \beta_1 + \beta_2 M_V + \beta_3 [\text{Fe}/\text{H}] + \beta_4 [\text{Fe}/\text{H}]^2 + \beta_5 [\text{Fe}/\text{H}] M_V \quad (3)$$

where M_V was chosen to be either $M_V(\text{BTO})$ or $M_V(\text{TO})$. The coefficients of the median and $\pm 1\sigma$ fits, for both $M_V(\text{BTO})$ and $M_V(\text{TO})$ are given in Table 5.

The use of these fits for determining ages via $M_V(\text{BTO})$ is illustrated for NGC 6752. Averaging the white dwarf distance modulus (§2.2) and the subdwarf main sequence fitting modulus (§2.3.1) results in $(m - M)_O = 13.12 \pm 0.07$, or $(m - M)_V = 13.25 \pm 0.07$. Using the photometry of Penny & Dickens 1986, we find $V(\text{BTO}) = 16.83 \pm 0.04$, so that $M_V(\text{BTO}) = 3.58 \pm 0.08$. Recall that $[\text{Fe}/\text{H}] = -1.51 \pm 0.08$ (§2.3.1). Using the coefficients of the fits in Table 1, this corresponds to an age of $11.15 \pm 0.8 \text{ Gyr}$ if $M_V(\text{BTO})$ and $[\text{Fe}/\text{H}]$ were known exactly. The effects of the $M_V(\text{BTO})$ and $[\text{Fe}/\text{H}]$ errors may be taken into account by using the median fit, and calculating ages for the $\pm 1\sigma$ values for $M_V(\text{BTO})$ and $[\text{Fe}/\text{H}]$. This procedure results in estimated errors of $\pm 0.9 \text{ Gyr}$ due to the $M_V(\text{BTO})$ error ($\pm 0.08 \text{ mag}$) and $\pm 0.4 \text{ Gyr}$ due to the $[\text{Fe}/\text{H}]$ error of $\pm 0.08 \text{ dex}$. Adding all three errors together in quadrature yields an age of $11.2 \pm 1.3 \text{ Gyr}$ for NGC 6752. This intermediate metallicity cluster has an age quite similar to the mean age of the 17 metal-poor clusters ($11.5 \pm 1.3 \text{ Gyr}$) determined in §4.2. More important, note that the uncertainty on the age of NGC 6752 determined in this way is comparable to the uncertainty in the mean of the set of 17 old globular clusters, illustrating the potential power of the method based on $M_V(\text{BTO})$.

Similarly, for M5, we calculate $V(\text{BTO}) = 18.03 \pm 0.02$ using the photometry of Sandquist *et al.* 1996. With $(m - M)_V = 14.51 \pm 0.09$ (§2.3.2), this results in $M_V(\text{BTO}) = 3.52 \pm 0.09$. Assuming $[\text{Fe}/\text{H}] = -1.17 \pm 0.08$ (Snedden *et al.* 1992), and using the technique outlined for NGC 6752, an age of $8.9 \pm 1.1 \text{ Gyr}$ is derived. Finally, for M13 with $[\text{Fe}/\text{H}] = -1.58$ and $(m - M)_V = 14.47 \pm 0.09$ (§2.3.3) we find $V(\text{BTO}) = 18.00 \pm 0.04$ using the photometry¹¹ of VandenBerg *et al.* 1990, resulting in $M_V(\text{BTO}) = 3.53 \pm 0.10$ and an age of $10.9 \pm 1.4 \text{ Gyr}$. Our results for the distances and ages of these three clusters are summarized in Table 6.

5. Summary

Our new work has two primary results. First, we have updated the absolute age estimate, and quantified the uncertainty in this estimate for the oldest globular cluster mean age. This update is primarily due to a reanalysis of estimates for the quantity which dominates the age uncertainty: $M_V(\text{RR})$ (the distance scale to GCs). We have concentrated on exploring in detail different estimates in order to account for the mean value, and the distribution in the uncertainty of this quantity. We find that all the data, not merely the recent Hipparcos parallax measurements, suggests a large systematic shift in $M_V(\text{RR})$ of approximately 0.2 magnitudes compared to earlier estimates. This has the effect of reducing the mean age of the oldest globular clusters by almost 3 Gyr. At the same time, this new data makes it clearer that $M_V(\text{RR})$ residual uncertainties are primarily systematic, reminding us that even apparently gaussianly distributed measurements in astrophysics may be subject to large systematic shifts. As a result, we now incorporate a large

¹¹The deep photometry of Richer & Fahlman 1986 used in the main sequence fitting (§2.3.3) contains very few subgiant stars, and so does not lead to a precise $V(\text{BTO})$ value. The photometry of VandenBerg *et al.* 1990 used here appears to be on the same system as the Richer & Fahlman 1986 photometry.

systematic uncertainty in the claimed mean value of $M_V(\text{RR})$ in our estimates.

Next, we provide a formalism which may be used by other researchers to update the estimates given here as new data emerges. In particular, we have presented an explicit discussion of the effect of other input parameter uncertainties from stellar evolution theory on the inferred GC ages estimates. We have displayed these effects in Figures 2-9, and provided analytical fits for both median ages, and uncertainties in age as a function of these parameters, and also as a function of $M_V(\text{RR})$.

We have also explicitly provided the fit for individual globular cluster ages and uncertainties as a function of metallicity and turn-off magnitude, using both the $M_V(\text{TO})$ and $M_V(\text{BTO})$ schemes. This should allow one to derive the age, and uncertainty in age for any GC with $-2.5 \leq [\text{Fe}/\text{H}] \leq -1.0$. We have illustrated this scheme, for the $M_V(\text{BTO})$ method for NGC 6752, using the average distance modulus from white dwarf sequence fitting, and subdwarf main sequence fitting, yielding an age of 11.2 ± 1.2 Gyr, illustrating that the $M_V(\text{BTO})$ method in principle allows an age precision on individual GC age determinations comparable to the $M_V(\text{TO})$ method applied to the ensemble of 17 old Globular clusters used in our analysis.

Finally, we briefly comment here on the cosmological implications of our central result that the mean of 17 old, metal-poor GC is 11.5 ± 1.3 Gyr, with a one-sided, 95% confidence level lower bound of 9.5 Gyr (see Krauss (1997) for further details). First and foremost, this results suggests that the long-standing conflict between the Hubble age, and GC age estimates for a flat matter dominated universe is now resolved for a realistic range of Hubble constants. A flat universe has an age which exceeds our lower limit on the GC ages for a Hubble constant $H_0 \leq 67 \text{ km s}^{-1} \text{ Mpc}^{-1}$, which is well within the range of current measured values. Thus, it now appears that the “age problem” is now no longer the primary motivation for considering a non-zero cosmological constant in the universe (i.e. Krauss and Turner 1995), and requires an alteration in the arguments associated with the debate between an open, flat matter dominated, and flat cosmological constant cosmologies (Krauss 1997).

If measurements of the Hubble constant continue to converge on the range $60\text{--}70 \text{ km s}^{-1} \text{ Mpc}^{-1}$, as suggested by the most recent analyses, cosmological concordance, at least as far as age is concerned, will perhaps for the first time be possible in all three scenarios.

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A. Potential Biases in the Subdwarf Sample

The parallaxes and absolute magnitudes for the stars listed in Table 2 do not include any statistical correction for possible biases in the sample. There are a few sources of potential biases in the sample. The classical Lutz & Kelker 1973 correction is a statistical correction which takes into account systematic effects due to the fact that (a) stars with parallaxes measured too high have a higher probability of being included in the sample than those with parallaxes measured too low (due to our $\sigma_\pi/\pi < 0.1$ selection criterion), and (b) more weight is given to stars with parallaxes that are overestimated rather than to stars with underestimated parallaxes (due to our use of a weighted least squares fit). In addition to this, Pont *et al.* 1997 point out that since metal-poor stars are far less numerous than more metal-rich stars, there may be an average underestimation of $[\text{Fe}/\text{H}]$ in the sample. The importance of these biases will depend on the selection criterion which are used to select the subdwarfs used in the main sequence fitting. The 3 papers which have used Hipparcos subdwarf parallaxes to determine GC distances have all had different selection criterion, and have determined different bias corrections. In their study, Pont *et al.* 1997 determined that the unevolved subdwarfs had a mean bias of +0.64 mag. In contrast, Gratton *et al.* 1997 determined a bias correction of -0.004 mag. Reid 1997 whose subdwarf sample consisted of high proper motion stars, elected to use individual Lutz-Kelker corrections, whose magnitude depended on the uncertainty in the parallax. In general, the corrections used by Reid 1997 were small, and in the opposite sense to those employed by Pont *et al.* 1997.

Our subdwarf study differs from the those of Reid 1997, Gratton *et al.* 1997 and Pont *et al.* 1997 in that we have access to the entire Hipparcos catalogue. Stars were selected for inclusion in the Hipparcos input catalogue based on a variety of considerations, and so there is no well defined selection criterion for the entire Hipparcos catalogue. Thus, it is difficult to assess the importance of the various biases *a priori*. For this reason, we have elected to use a stringent selection criterion $\sigma_\pi/\pi < 0.1$ which minimizes the importance of the Lutz & Kelker 1973 type bias (Brown *et al.* 1997). As it turns out, the final sample only contains stars with $\sigma_\pi/\pi < 0.08$, strongly suggesting that the stars whose true parallax are systematically smaller than the observed parallax are not preferentially included in our sample.

To study the possible biases which remain in our subdwarf sample, we have constructed a Monte Carlo simulation to generate synthetic data whose properties are known, and compared to ‘observed’ properties which are calculated in the Monte Carlo. This is similar in spirit to the bias studies of Gratton *et al.* 1997 and Pont *et al.* 1997. We have attempted to construct a subdwarf data set whose properties and selection biases closely match those in our actual data set. In particular, our subdwarf sample consists of stars with $[\text{Fe}/\text{H}] < -1.0$, $M_V > 5.5$, and $\sigma_\pi/\pi < 0.1$ and these facts are incorporated in the Monte Carlo. The Monte Carlo was constructed in the following steps:

1. An intrinsic $[\text{Fe}/\text{H}]$ value (below $[\text{Fe}/\text{H}] = -1.0$) was drawn from one of two probability functions. The first function is that given by the observed $[\text{Fe}/\text{H}]$ distribution in the Carney

et al. 1994 study

$$P([\text{Fe}/\text{H}]) = 47.13 + 14.35[\text{Fe}/\text{H}]. \quad (\text{A1})$$

This function was chosen as many of the metal-poor stars in the Hipparcos input catalog are in the Carney *et al.* 1994 [Fe/H] catalogue. The second [Fe/H] distribution we considered was that given by Pont *et al.* 1997

$$P([\text{Fe}/\text{H}]) = 1.4 \exp([\text{Fe}/\text{H}] + 3) - 1 \quad (\text{A2})$$

and represents their approximation to the observed [Fe/H] distribution in their sample.

2. An observed [Fe/H] value $[\text{Fe}/\text{H}]_o$ was determined from the intrinsic [Fe/H] value by adding a random value which was taken from gaussian distribution with $\sigma_{[\text{Fe}/\text{H}]} = 0.10, 0.15$ and 0.20 .
3. The distance d was determined assuming a sphere of uniform density and the true parallax was determined, $\pi_t = 1/d$.
4. The absolute magnitude (M_V) was determined assuming an Salpeter initial mass function ($\Phi(m) \propto m^{-2.35}$, with upper and lower mass limits taken to be $m = 0.9$ and $0.4 M_\odot$) and a mass luminosity relation taken from our standard isochrones:

$$\begin{aligned} M_V &= 13.81 - 12.11m, \quad \text{for } [\text{Fe}/\text{H}] < -1.5 \text{ and} \\ M_V &= 15.14 - 13.02m \quad \text{for } -1.0 \leq [\text{Fe}/\text{H}] \leq -1.5. \end{aligned} \quad (\text{A3})$$

5. The true apparent magnitude V_t was calculated from the absolute magnitude and true parallax $V_t = M_V - 5.0 \log(\pi_t) - 5.0$.
6. The observed magnitude V_o was calculated from the apparent magnitude assuming gaussian errors with $\sigma_V = 0.02$
7. To reproduce the Hipparcos catalogue completeness characteristics, we assumed the catalogue was complete up to $V_o = 9$ and increasingly incomplete fainter than this, with a probability for inclusion of

$$P(V) = 10^{-\tau(V_o - 9)} \quad (\text{A4})$$

where τ was chosen to be 0.6 which is valid for the entire Hipparcos catalogue (Pont *et al.* 1997), or $\tau = 0.15$ which is valid for the metal-poor stars in the Carney *et al.* 1994 [Fe/H] catalogue.

8. The observed parallax (π_o) was computed from the true parallax assuming gaussian errors with σ_π . The value of σ_π was taken from a fit to the parallax errors in the Hipparcos catalogue

$$\sigma_\pi = -3.96 + 1.893V_o - 0.26465 * V_o^2 + 0.013107V_o^3 \quad (\text{A5})$$

in units of mas. The scatter of the parallax errors about the mean value given by equation (A5) was taken into account by adding a random value to σ_π which was taken from gaussian

distribution with a σ varying from 0.54 to 1.14 mas between $V_o = 5$ to 12. This derived parallax error will be referred to as σ_{π_o} . This procedure accurately reproduces the observed parallax errors as a function of apparent magnitude found in Hipparcos catalogue.

9. The observed absolute magnitude M_{V_o} was calculated

$$M_{V_o} = V_o + 5.0 \log(\pi_o) + 5.0. \quad (\text{A6})$$

10. The simulated data are accepted if (a) $\sigma_{\pi_o}/\pi_o < 0.10$, (b) $M_{V_o} > 5.5$ and (c) $[\text{Fe}/\text{H}]_o$ was below some value. The $[\text{Fe}/\text{H}]$ cutoff was allowed to vary, so that some runs required $[\text{Fe}/\text{H}]_o < -1.0$ (valid for our sample), and others required $[\text{Fe}/\text{H}]_o < -1.8$, the cutoff used by Pont *et al.* 1997

A typical simulation contained 10^7 simulated input stars, of which $\sim 10^4$ were accepted. For the data which was accepted, the mean absolute magnitude and $[\text{Fe}/\text{H}]$ biases were calculated. Both the weighted mean, and unweighted mean bias was calculated. As we use weighted fits in the subdwarf fitting analysis, it is the weighted mean bias which is appropriate for our sample. However, Pont *et al.* 1997 determined an unweighted mean bias, so this quantity was calculated as well in order to compare our results to Pont *et al.* 1997. The results are summarized in Table 7 for the various cases given above. In all cases, we found that the weighted mean absolute magnitude bias was small. The largest (in absolute value) absolute magnitude weighted bias was -0.006 mag. This translates into an age reduction of less than 0.1 Gyr. Given the small value of this correction, we have elected not applied it to our subdwarf fitting. Our results are in good agreement with those of Gratton *et al.* 1997. The unweighted mean absolute magnitude bias is typically a factor of 10 larger than the weighted mean, but is still relatively small (maximum absolute value of -0.034 mag). In no case did we find a positive absolute magnitude bias (as was found by Pont *et al.* 1997).

The weighted mean $[\text{Fe}/\text{H}]$ bias could be as large as $+0.12$ dex for stars selected to have $[\text{Fe}/\text{H}] < -1.8$. However, we did not use any stars with $[\text{Fe}/\text{H}] < -1.8$ in our main sequence fitting analysis. Considering the samples with have a metallicity cut at $[\text{Fe}/\text{H}] < -1.0$, the weighted mean $[\text{Fe}/\text{H}]$ bias is likely to lie in the range $+0.01$ to $+0.08$ dex. We believe case B best represents the true subdwarf distribution; it has a weighted mean $[\text{Fe}/\text{H}]$ bias of $+0.03$ dex, which (considering the results presented in §2.3.1 to 2.3.3) could result in an absolute magnitude bias correction up to $+0.015$ mag. Given the small value of this correction, and the fact that it acts in the opposite sense to the absolute magnitude bias correction determined above, we have elected not to apply it to our main sequence fitting results. Our main sequence fitting results allow for up to a $+0.20$ dex systematic error in the subdwarf metallicity scale.

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Table 1. Astrometric Distances

Cluster	[Fe/H]	(m – M) _O	V(HB)	M _V (HB)	[Fe/H] reference
M5	–1.17	14.44	15.05	0.51 ± 0.41	Snedden <i>et al.</i> 1992
M4	–1.33	11.18	13.37	0.67 ± 0.23	Zinn & West 1984
M22	–1.75	12.17	14.10	0.58 ± 0.19	Zinn & West 1984
M3 ^a	–1.47	14.91	15.63	0.69 ± 0.59	Kraft <i>et al.</i> 1992
M13 ^b	–1.58	14.06	14.83	0.71 ± 0.23	Kraft <i>et al.</i> 1997
M92	–2.25	14.76	15.13	0.31 ± 0.32	Snedden <i>et al.</i> 1991

^aV(HB) from Buonanno *et al.* 1994a. Adopted reddening of 0.01 from Zinn 1985.

^bV(HB) from Buonanno *et al.* 1989

Table 2. Calibrating Subdwarfs

Name	V _O	(B – V) _O	E(B – V)	[Fe/H]	π (mas)	σ_{π}/π	M _V
HD 193901	8.65	0.56	0.00	–1.07	22.88	0.054	5.45 ± 0.117
HD 145417	7.53	0.82	0.00	–1.15	72.75	0.011	6.84 ± 0.024
HD 103095	6.43	0.75	0.00	–1.22	109.21	0.007	6.62 ± 0.015
HD 120559	7.97	0.66	0.00	–1.23	40.02	0.025	5.98 ± 0.054
HD 126681	9.28	0.61	0.00	–1.45	19.16	0.075	5.69 ± 0.163
BD+59 2407	10.20	0.58	0.05	–1.75	15.20	0.080	6.11 ± 0.174
HD 25329	8.51	0.87	0.00	–1.76	54.14	0.020	7.18 ± 0.043
CPD-80 349	10.05	0.54	0.02	–2.26	16.46	0.060	6.13 ± 0.130

Table 3. $M_V(\text{RR})$ Calibration

Method	[Fe/H]	$M_V(\text{RR})$	$M_V(\text{RR})$ at [Fe/H] = -1.9
Astrometric	-1.59	0.59 ± 0.11	0.52 ± 0.11
White dwarf fitting to N6752	-1.51	0.45 ± 0.14	0.36 ± 0.14
Subdwarf fitting to N6752	-1.51	0.30 ± 0.15	0.21 ± 0.15
Subdwarf fitting to M5	-1.17	0.54 ± 0.09	0.37 ± 0.09
Subdwarf fitting to M13	-1.58	0.36 ± 0.14	0.29 ± 0.14
LMC RR Lyr	-1.90	0.44 ± 0.14	0.44 ± 0.14
Theoretical models	-2.20	0.36 ± 0.10	0.43 ± 0.10

Table 4. Monte Carlo Input Parameters

Parameter	Distribution	Comment
mixing length	1.85 ± 0.25 (stat.)	fits GC observations
helium diffusion coefficients	$0.3 - 1.2$ (syst.)	possible systematic error dominate
high temperature opacities	1 ± 0.01 (stat.)	comparison of OPAL & LAOL opacities
low temperature opacities	$0.7 - 1.3$ (syst.)	comparison of different tables
primordial ^4He abundance	$0.22 - 0.25$ (syst.)	possible systematic error dominate
oxygen abundance, $[\text{O}/\text{Fe}]$	$+0.55 \pm 0.05$ (stat.) ± 0.20 (syst.)	mean from Nissen <i>et al.</i> 1994
surface boundary condition	grey or Krishna-Swamy 1966	
colour table	Green <i>et al.</i> 1987 or Kurucz 1992	
Nuclear Reaction Rates:		
$p + p \longrightarrow ^2\text{H} + e^+ + \nu_e$	1 ± 0.002 (stat.) $+0.0014 \ +0.02$ $-0.0009 \ -0.012$ (syst.)	see Paper I
$^3\text{He} + ^3\text{He} \longrightarrow ^4\text{He} + 2p$	1 ± 0.06 (stat.)	Bahcall & Pinsonneault 1992
$^3\text{He} + ^4\text{He} \longrightarrow ^7\text{Be} + \gamma$	1 ± 0.032 (stat.)	Bahcall & Pinsonneault 1992
$^{12}\text{C} + p \longrightarrow ^{13}\text{N} + \gamma$	1 ± 0.15 (stat.)	Bahcall 1989, table 3.4
$^{13}\text{C} + p \longrightarrow ^{14}\text{N} + \gamma$	1 ± 0.15 (stat.)	Bahcall 1989, table 3.4
$^{14}\text{N} + p \longrightarrow ^{15}\text{O} + \gamma$	1 ± 0.12 (stat.)	Bahcall 1989, table 3.4
$^{16}\text{O} + p \longrightarrow ^{17}\text{F} + \gamma$	1 ± 0.16 (stat.)	Bahcall 1989, table 3.4

Table 5. Fit coefficients for age as a function of $M_v(\text{BTO})$ and $M_v(\text{TO})$

	β_1	β_2	β_3	β_4	β_5
$M_v(\text{BTO})$ median	-0.824	0.418	-0.248	-0.033	-0.014
$M_v(\text{BTO}) +1\sigma$	-0.775	0.418	-0.221	-0.030	-0.017
$M_v(\text{BTO}) -1\sigma$	-0.857	0.413	-0.266	-0.037	-0.014
$M_v(\text{TO})$ median	-1.305	0.515	-0.396	-0.018	0.049
$M_v(\text{TO}) +1\sigma$	-1.322	0.524	-0.428	-0.024	0.052
$M_v(\text{TO}) -1\sigma$	-1.285	0.505	-0.361	-0.011	0.044

Table 6. GC distances and Ages

Name	[Fe/H]	E(B – V)	(m – M) _V	Age (Gyr)
NGC 6752	-1.51 ± 0.08	0.04 ± 0.01	13.25 ± 0.07	11.2 ± 1.3
M5	-1.17 ± 0.08	0.03 ± 0.01	14.51 ± 0.09	8.9 ± 1.1
M13	-1.58 ± 0.08	0.02 ± 0.01	14.47 ± 0.09	10.9 ± 1.4

Table 7. Monte Carlo Subdwarf Bias Results

Case	[Fe/H] distribution	$\sigma_{[\text{Fe}/\text{H}]}$	τ	[Fe/H] cut	weighted mean bias		unweighted mean bias	
					M _V (mag)	[Fe/H] (dex)	M _V (mag)	[Fe/H] (dex)
A	Carney <i>et al.</i> 1994	0.10	0.15	< -1.0	–0.003	+0.011	–0.034	+0.004
B	Carney <i>et al.</i> 1994	0.15	0.15	< -1.0	–0.004	+0.033	–0.034	+0.009
C	Carney <i>et al.</i> 1994	0.20	0.15	< -1.0	–0.005	+0.057	–0.034	+0.016
D	Carney <i>et al.</i> 1994	0.10	0.60	< -1.0	–0.003	+0.022	–0.025	+0.003
E	Carney <i>et al.</i> 1994	0.20	0.60	< -1.0	–0.004	+0.064	–0.026	+0.013
F	Pont <i>et al.</i> 1997	0.15	0.15	< -1.0	–0.004	+0.033	–0.034	+0.009
G	Pont <i>et al.</i> 1997	0.20	0.15	< -1.0	–0.006	+0.074	–0.031	+0.025
H	Pont <i>et al.</i> 1997	0.20	0.60	< -1.0	–0.005	+0.077	–0.024	+0.019
I	Pont <i>et al.</i> 1997	0.15	0.60	< -1.8	–0.004	+0.076	–0.028	+0.047
J	Pont <i>et al.</i> 1997	0.20	0.60	< -1.8	–0.005	+0.125	–0.028	+0.077

Fig. 1.— Unevolved ($M_V \gtrsim 5.5$), metal-poor ($[\text{Fe}/\text{H}] \lesssim -1.0$) stars in the Hipparcos catalogue which have very good parallaxes ($\sigma_\pi/\pi < 0.10$) and which are not known binaries are compared to our isochrones. Each Hipparcos star (points with error bars) is labeled with its spectroscopic $[\text{Fe}/\text{H}]$ value.

Fig. 2.— Histogram of GC ages. The median, mean, standard deviation and one-sided, 95% confidence level lower limit are all indicated on the figure.

Fig. 3.— Best estimate for the age of the oldest GCs, as a function of the adopted $M_V(\text{RR})$ value at $[\text{Fe}/\text{H}] = -1.9$. The best fitting median (solid line), along with $\pm 1\sigma$ limits (dashed lines) are plotted. These lines are log fits ($\log(t_9) = a + b M_V(\text{RR})$), with the following coefficients: median $(a, b) = (0.888, 0.454)$; $-1\sigma(a, b) = (0.866, 0.436)$; and $+1\sigma(a, b) = (0.915, 0.463)$. Further details are provided in the text.

Fig. 4.— Age as a function of $[\alpha/\text{Fe}]$. The best fitting median, along with $\pm 1\sigma$ limits are plotted. These lines are of the form $t_9 = a + b[\alpha/\text{Fe}]$, with the following coefficients: median $(a, b) = (13.83, -3.77)$; $-1\sigma (a, b) = (13.26, -3.72)$; and $+1\sigma (a, b) = (14.54, -4.00)$.

Fig. 5.— Age as a function of the mixing length (α) used in the stellar models. The lines of the form $t_9 = a + b\alpha$ have the following coefficients: median $(a, b) = (9.66, 1.13)$; $-1\sigma(a, b) = (9.06, 1.10)$; and $+1\sigma(a, b) = (10.31, 1.17)$.

Fig. 6.— Age as a function of the helium abundance (Y) used in the stellar models. The lines of the form $t_9 = a + bY$ have the following coefficients: median $(a, b) = (19.05, -31.15)$; $-1\sigma (a, b) = (17.61, -27.73)$; and $+1\sigma (a, b) = (18.80, -27.04)$.

Fig. 7.— Age as a function of the $^{14}\text{N} + \text{p} \rightarrow ^{15}\text{O} + \gamma$ reaction rate ($\mathfrak{R}$). The lines of the form $t_9 = a + b \mathfrak{R}$ have the following coefficients: median $(a, b) = (12.97, -1.26)$; $-1 \sigma (a, b) = (12.09, -1.01)$; and $+1 \sigma (a, b) = (14.22, -1.76)$.

Fig. 8.— Age as a function of the helium diffusion coefficient (D) used in the stellar models. The lines of the form $t_9 = a + b D$ have the following coefficients: median $(a, b) = (12.23, -0.67)$; $-1\sigma (a, b) = (11.61, -0.71)$; and $+1\sigma (a, b) = (13.05, -0.76)$.

Fig. 9.— Age as a function of low temperature opacity (κ). The best fitting median, along with $\pm 1\sigma$ limits are plotted. These lines are of the form $t_9 = a + b\kappa$, with the following coefficients: median $(a, b) = (12.41, -0.68)$; $-1\sigma (a, b) = (11.63, -0.55)$; and $+1\sigma (a, b) = (13.16, -0.67)$.

Fig. 10.— Histograms for the mean age of the oldest globular clusters, using (a) the RYI (Green *et al.* 1987) colour table, or (b) the Kurucz 1992 colour table.